Trilemma Stability and International Macroeconomic Archetypes

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Abstract

This paper uses the simple geometry of the classic, open-economy trilemma to introduce a new gauge of the stability of international macroeconomic arrangements. The new stability gauge reflects the simultaneity of a country's choices of exchange rate fixity, financial openness, and monetary sovereignty. So, the new gauge is bounded and correspondingly non-Gaussian. We use the new stability gauge in nonlinear panel estimates to examine the post-Bretton Woods period, and we find that trilemma policy stability is linked to official holdings of foreign exchange reserves in low income countries. We also find that the combination of fixed exchange rates and financial market openness is the most stable arrangement within the trilemma; and middle-income countries have less stable trilemma arrangements than either

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low or high-income countries. The paper also characterizes international macroeconomic arrangements in terms of their semblance to definitive policy archetypes; and, it uses the trilemma constraint to provide a new gauge of monetary sovereignty.

Keywords: Trilemma, Foreign Exchange Rate Regimes, Exchange Rates, International Reserves, Financial Openness, Fear of Floating, Monetary Sovereignty

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1. Introduction

The classic, open-economy trilemma tells us that a country cannot simultaneously achieve exchange rate stability, capital market openness, and monetary sovereignty. Choosing, say, to peg an exchange rate means choosing to give up some degree of monetary sovereignty, capital market openness, or both. While the trilemma demands that such choices be made, the choices are never final.[1] This paper introduces a new, formal measure of the stability – or instability – of such arrangements. Based on the constraints of the trilemma itself, the new measure is bounded and drawn from a non-Gaussian distribution. As measured here, trilemma policy changes are thus themselves non-normal. This paper uses the new measure to describe the incidence of policy changes during the post-Bretton Woods period, and it explores the policy changes further using nonlinear panel estimates.

[1] That exchange rate arrangements are not permanent has been highlighted by recent history, emphasized by Obstfeld and Rogoff (1995), and further explored in Calvo and Reinhart (2002), Reinhart and Rogoff (2004), Levy-Yeyati and Sturzeneger (2005), and Ilzetzki et al. (2011). Those papers, among others, document the sometimes dramatic changes in *de facto* exchange rate regimes. In this paper, we build on such studies of exchange rate stability by encompassing all three legs of the trilemma, rather than just the exchange rate itself.

The new measure of stability starts with the simple geometry of the trilemma. We can think of a country's international macroeconomic arrangements in terms of locations in a constrained three-dimensional policy space, one that is defined by exchange rate stability, financial openness, and monetary policy sovereignty. In this framework, the *change* in a country's arrangement is naturally measured as a movement from one point to another in the three-dimensional policy space. So, the stability or instability of a country's arrangements is reflected in the extent of the changes over time: it is measured by the distances between the sequential locations in the policy space. A stable arrangement is defined as one with relatively small movements within the policy space, while large movements within the policy space represent unstable arrangements.

We also provide a new measure of monetary sovereignty. While there are several existing approaches to measuring capital mobility and exchange rate policy, that is not the case for monetary sovereignty. The extant literature has only one well-used approach to measuring sovereignty. That approach relies on the correlation between a country's interest rate and the interest rate of a base country. One drawback to using such correlations is that they often conflate monetary dependence with other sources of shared dynamics. The new measure presented here does not use interest rate correlations. Instead, it is derived from the trilemma's constraint. The trilemma constrains monetary sovereignty at the expense of exchange rate stability and financial openness. So, given measures of exchange rate stability and financial openness, the trilemma's constraint implicitly provides a measure of monetary sovereignty. This new, implicit measure complements the now-standard interest rate correlation approach.

We use the trilemma stability and monetary sovereignty measures to ex-

plore which types of trilemma policies are most stable and to study whether official foreign exchange reserves are related to greater trilemma stability. In the next section of this paper, we introduce our new measures, first of stability then of monetary sovereignty. We then use the measures to assess the stability of the trilemma policies in the modern era. Next, we sort countries into policy archetypes in each year and explore the stability of the archetypes. Finally, we examine the links between stability, archetype, and official holdings of foreign reserves.

2. Two New Measures

2.1. A Stability Measure

To gauge stability, we begin with the international trilemma's standard triad of policies. We denote the i^{th} country's extant regime in period t as $R_{i,t}$, where:

$$R_{i,t} = (S_{i,t}, F_{i,t}, M_{i,t}),$$

and $S_{i,t}$ represents exchange rate stability, $F_{i,t}$ represents financial openness, and $M_{i,t}$ represents monetary sovereignty. The measures of $S_{i,t}$, $F_{i,t}$, and $M_{i,t}$, are normalized so that each falls between zero and one (inclusive); and values of one represent perfectly fixed exchange rates, perfectly open financial markets, and perfectly sovereign monetary policy. So, a pure fix with open financial markets is: $R_{i,t} = (1, 1, 0)$; a pure fix with monetary sovereignty is $R_{i,t} = (1, 0, 1)$, and a pure float with open capital markets and monetary sovereignty is $R_{i,t} = (0, 1, 1)$.

In this framework, a *change* in a country's regime from one period to the next is simply the vector connecting the two consecutive points in the

policy space:

$$r_{i,t} = R_{i,t} - R_{i,t-1} = (s_{i,t}, f_{i,t}, m_{i,t}) = (S_{i,t} - S_{i,t-1}, F_{i,t} - F_{i,t-1}, M_{i,t} - M_{i,t-1}).$$

Using this vector of policy changes, $r_{i,t}$, we can definitively measure the overall change in policy using the vector's norm, $||r_{i,t}||$.[2] Using the norm, we define a single, univariate measure adjusted to fall between zero and one:

$$n_{i,t} = \frac{||r_{i,t}||}{\sqrt{2}}.$$

This adjusted norm, $n_{i,t}$, captures in a simple scalar the full extent of the change in a country's triad of policies. A value of $n_{i,t}$ equal to zero would mean that a country has not changed its three policies since the previous year. In contrast, a large value of $n_{i,t}$ would reflect a substantial change relative to the prior year.

By reducing three dimensions to one, the norm gauges the stability between periods of the triad of policies within the trilemma. That said, the measure has two potential conceptual drawbacks. First, it requires that we make an assumption about the functional form of the trade-offs between policies. While most open-economy macroeconomic models implicitly include the trilemma as an arbitrage-like condition, various models differ in terms of the functional forms they would imply for the trade-offs among the

[2]We use the familiar l^2-norm, or *Euclidean* norm. That is, we use: $||r_{i,t}|| = (s_{i,t}^p + f_{i,t}^p + m_{i,t}^p)^{\frac{1}{p}}$, with $p = 2$. However, we also calculate the *taxi* norm, $p = 1$, and the *infinity* norm, $p = \infty$. Despite their different intuitive interpretations (the Euclidean norm is the distance 'as the crow flies,' the taxi norm adds up the full change in each dimension, and the infinity norm takes the largest move in any of the dimensions), the kernel densities of these norms are similarly shaped, and the full sample estimates reported in the panel results below are not sensitive to the use of these alternative norms. We refer to the the l^2-norm as "the" norm in the rest of the paper.

policies. In this paper, we assume that the trilemma constraint is a linear one in the normalized units that we adopt.[3] This assumption has the virtue of simplicity, and it is supported empirically by Aizenman et al. (2008) and Wu (2011). The second potential drawback is that, while the norm provides a gauge of policy stability that reflects the idea that no single policy can be changed on its own, it does not, by itself, retain information about which of the two or three policies has changed. This second drawback can be addressed by using the norm in conjunction with other data. For example, in the empirical work below, we combine the observations of the norm with observations of the trilemma's individual pieces.

By providing a univariate gauge of multivariate changes in policies, our new measure reflects the spirit of the Girton and Roper (1977) 'exchange market pressure' measure. Their measure provided an early, univariate amalgam of foreign exchange policies. Our measure is a similar amalgam, one that now has a geometric interpretation within the well-known trilemma.

[INSERT FIGURE 1 ABOUT HERE]

Figure 1 illustrates our approach to measuring policy stability. The figure displays the two data points underlying a single observation of the adjusted norm, $n_{i,t}$.[4] The observation is for Indonesia at the time of the Asian Crisis (i = Indonesia, and $t = 1997$), and the underlying data are from Aizenman, Chinn, and Ito (2010), which we discuss in more detail in section 3.1.[5] As is well-known, Indonesia experienced a substantial increase in its exchange rate

[3] That is, the trilemma can be viewed as a triangular surface in three dimensional space, as illustrated in Appendix Figure A1.

[4] Note that the linear version of the trilemma constraint would require that all points lie on the plane defined by the three points: $(0, 1, 1), (1, 0, 1)$, and $(1, 1, 0)$.

[5] This figure uses our new, implicit measure of monetary sovereignty, also described below.

variability and a small reduction in its financial openness during the crisis, while it increased its monetary sovereignty considerably. These changes are indicated in the figure by the vector shown between the observations for 1996 and for 1997.[6] The normalized length of the vector measures the overall change in the policy triad. The norm in 1997 is about five times the values typical of Indonesia earlier in the decade, and it exceeds (by a substantial margin) 95 percent of the values in the sample. After introducing the data, in section 3, we provide additional figures and summary statistics.

In general, the norm of the vector summarizes the overall changes in the policies of the trilemma. Below, we use the norm (adjusted to fall between zero and one) to examine the stability of various policies and to assess the extent to which stability may be linked to official holdings of foreign exchange reserves.

2.2. An Implicit Measure of Monetary Sovereignty

The most often-used measure of monetary sovereignty relies on the approach of Shambaugh (2004). That approach reflects the correlation between a country's domestic, short-term interest rate and that of a putative base country, often the United States. High correlations are taken as indicative of monetary dependence. That is, they are taken as a lack of monetary sovereignty. The drawback of this otherwise valuable approach is that, in addition to monetary dependence, the measure also captures the interest rate effects of the underlying circumstances to which independent monetary policies may or may not respond. So, at one extreme, even a country with complete monetary sovereignty appears otherwise when it is subject to some

[6] The cartesian coordinates $(S_{i,t}, F_{i,t}, M_{i,t})$ are (0.66, 0.94, 0.4) for 1996 and (0.11, 0.88, 1.0) for 1997. So, $n_{i,t} = 0.578$.

of the same shocks or influences as its putative base country. At the other extreme, a country with no monetary sovereignty might misleadingly appear to be quite autonomous when it is subject to disturbances not experienced by its base country.

New Zealand provides a telling example of the standard measure's problem. The Reserve Bank of New Zealand is the prototypic inflation targeter. While it could conceivably be influenced by the policies of Australia (its "base country" in Shambaugh's work), it is in no way constrained by Australia's policies. Nevertheless, the interest rates of New Zealand and Australia are – as one might expect – often highly correlated. So, taken at face value, the standard approach might wrongly seem to suggest that New Zealand's monetary policy is dictated by the Reserve Bank of Australia.

Other researchers, such as Frankel et al. (2004), and Reade and Volz (2010), allow for more general dynamic links between the interest rates of the countries. However, even these more general measures ultimately rely on interest rate comovements, so they are subject to the same drawback.[7]

Here, we introduce an alternative measure of monetary sovereignty that does not suffer from this drawback, and we use the new measure of sovereignty in our gauge of stability, $n_{i,t}$. Our new measure of sovereignty starts from the trilemma itself. Specifically, we maintain our assumption that the trilemma holds linearly. With that assumption, the existing measures of exchange rate stability, $S_{i,t}$, and of financial openness, $F_{i,t}$, provide us with a very simple,

[7]Three other, more recent studies take important steps toward mitigating the problem. Duburcq and Girardin (2010) allow domestic monetary conditions to matter in a study of eight Latin American countries over eleven years. Bluedorn and Bowdler (2010) separate the anticipated and unanticipated components of the base country's interest rate changes using the U.S. as the base country. Herwartz and Roestel (2010) examine long-run interest rate dependence and condition on domestic variables for a panel of 20 small, high income countries.

implicit measure of monetary sovereignty, $M_{i,t}$. Specifically, the implicit measure of monetary sovereignty is:

$$M_{i,t} = 2 - S_{i,t} - F_{i,t}.$$

Using data from Aizenman et al. (2010), described in more detail below, Figure 2 depicts both this new measure (the blue lines) and the interest rate correlation measure (the red lines).[8] Looking at the means, shown in the first row, the new, implicit measure suggests a greater degree of monetary sovereignty than does the interest rate measure. However, underlying these means are individual instances of differences in both directions, as show in the figure's middle row.

[INSERT FIGURE 2 ABOUT HERE]

For some economies, especially for those that peg exchange rates or maintain them in a narrow band, the new measure often indicates that there is *less* sovereignty than would be suggested by the interest rate correlation measure. This is the case for Hong Kong, shown in the row's first chart. The Hong Kong Monetary Authority tightly controls the value of exchange rate, and capital is allowed to move into and out of its economy relatively freely.[9] The trilemma tells us that in such cases there is little scope for monetary sovereignty, and in 2010 (the latest year in our sample), the new, implicit measure of sovereignty equals zero. In contrast, differences in the

[8] In cases where the implicit measure would yield a value in excess of one, we have equated the measure with one. The imposition of this limit reflects the fact that countries not pursuing exchange rate stability and financial openness to the fullest extent nevertheless cannot acquire more than complete ($M_{i,t} = 1$) monetary sovereignty.

[9] The Hong Kong Monetary Authority has assiduously pegged the Hong Kong dollar to the U.S. dollar since the eighties, and the United States is the base country in Shambaugh (2004).

behavior of U.S. and Hong Kong interest rates at times give rise to much higher correlation-based measures of sovereignty despite the tight peg. Hong Kong's correlation-based measure for 2010 is 0.45, a value that would seem to suggest that Hong Kong retained a good deal of monetary sovereignty, more so even than Australia.[10] Throughout all of Hong Kong's peg, the new, trilemma-implied sovereignty measures indicate that Hong Kong's monetary sovereignty was more limited than the interest rate correlations would have suggested.

For still other countries, the two sovereignty measures are quite similar; and, the measures occasionally are even identical. For example, both monetary sovereignty measures assign values of zero to eurozone economies in recent years, as illustrated by Austria, shown in the row's next chart.[11]

In some cases, the the new, implicit measure is much larger than the existing, correlation-based measure. For example, returning to the case of New Zealand, shown next, the 2010 interest rate correlation measure is only 0.17, a value that would seem to suggest that the Reserve Bank of New Zealand follows the monetary policies of Australia. In contrast, New Zealand's new, trilemma-based measure is much higher, 0.71, which reflects its substantial degree of monetary sovereignty. Similarly, Canada's monetary sovereignty (now also used to target inflation) is largely masked by the interest rate measure, which remains low as long as Canadian and U.S. interest rates continue to be relatively highly correlated. Canada's measures are shown in the row's last chart. For 2010, Canada's interest rate correlation-based measure is a modest 0.29, while the implied trilemma

[10] Australia's 2010 correlation-based measure was 0.37.

[11] In this chart, one can also see the onset of Austria's informal monetary union with Germany in 1981.

measure is 0.72.

Overall, the sample correlation between the trilemma-based sovereignty measure and the interest rate correlation-based sovereignty measure is 0.37. The sample correlation between the two measures is higher for high income countries, where it equals 0.53. It is lowest, 0.15, for middle income countries; and it is 0.21 for low income countries.

The final row of Figure 2 shows the average *changes* in the two monetary sovereignty measures. The blue lines depict the changes in the new, implicit measure; and the red lines depict the changes in the interest rate correlation measure. Using the new monetary sovereignty measure, it is now easy to see the monetary upheaval many countries (especially the high income ones, shown in the lower left) experienced in the wake of the Bretton Woods breakdown. By comparison, the correlation-based measures would have suggested that the rich countries experienced only modest changes in their monetary sovereignty. In more recent years, we see that changes in the correlation-based measures suggest a loss in sovereignty among low and middle income economies that does not appear so striking in the new measure.

As might be expected, the sample correlation between the changes in the two sovereignty measures is much smaller (0.04) than the sample correlation of their levels. The pattern across the income groups, however, remains the same. At 0.08, the correlation between the changes in the two measures is highest for the high-income group. The middle-income group has the lowest correlation, 0.02; while the correlation in the low-income group equals the average, 0.04.

Overall, while there are some exceptions (such as the Bretton Woods breakdown), both the standard deviations in Table 1 and the plots in Figure 2 suggest that the new, implicit measure is somewhat less variable than the

old one. That is, using the new, implicit measures, the greater relative sovereignty is ccompanied by a greater steadiness as well.

3. Data and Overall Trilemma Stability

3.1. Data Definition and Descriptive Statistics

In this section, we calculate the new measures of trilemma stability using a sample of 177 economies with annual data from 1970 through 2010. We begin with the data provided in Aizenman et al. (2010), updated with the latest version of the *de jure* financial account openness measure of Chinn and Ito (2006). Then, we recalculate our measure of trilemma stability using our new, implicit gauge of monetary sovereignty.

Aizenman et al. (2010) construct the annual measure of $S_{i,t}$, using the exchange rate's monthly standard deviation against a base country.[12,13] Like many other researchers, they follow Shambaugh (2004) in constructing monetary sovereignty measures, $M_{i,t}$, using the correlation between each country's money market interest rate and that of its base country. Their measure of financial market openness, $F_{i,t}$, is a *de jure* one: essentially, it is a weighted

[12] Aizenman et al. (2010) provide a continuous measure of $S_{i,t}$ that does not rely on the use of reserves to categorize exchange rate regimes. Other prominent *de facto* measures of exchange rate arrangements include: Shambaugh (2004), and later Klein and Shambaugh (2008), who classify exchange rate arrangements into floating and non-floating; Reinhart and Rogoff (2004), and more recently Ilzetzki et al. (2011), who rely on exchange rate behavior and more nuanced assessments to construct five coarse and many finer categories; and Levy-Yeyati and Sturzeneger (2005), who use information about exchange rates and about reserves to provide a cluster-based exchange rate taxonomy.

[13] Like others, Aizenman et al. (2010) apply a threshold to the standard deviation method in order to allow for currencies that remain in narrow bands; and, they also allow for individual devaluations or revaluations. The base countries include Australia, Belgium, France, Germany, India, Malaysia, South Africa, the United Kingdom, and the United States.

average of the International Monetary Fund's indicators of exchange restrictions.[14]

Table 1 provides summary statistics for the adjusted norms, $n_{i,t}$, calculated using these data and reported separately for each of the two sovereignty measures. Note that it is the stability of *policy* that is the focus here, not the stability of the exchange rate. In particular, a sustained float – with its inherent exchange rate volatility – can be part of a stable *policy*.[15] The first panel reports the statistics by income group, while the second panel reports them by decade.[16] The third panel provides the measures for the policy archetypes that are described later in this section. Finally, the bottom panel provides the summary statistics for the sample as a whole.

[INSERT TABLE 1 ABOUT HERE]

The first two columns of numbers report the mean and median for each category. In all cases, and using both measures, the means exceed the medians.

[14]Specifically, Chinn and Ito (2006) measure financial openness with the first principal component of the IMF's binary indicators of restrictions on current and capital account transactions, of multiple exchange rates, and of the required surrender of export proceeds. This is also the measure subsequently used by Aizenman et al. (2010). Miniane (2004) provides a *de jure* index that uses finer IMF data on capital account restrictions, but the data are available for only thirty countries. Many other, related, *de jure* indices have been developed, but few blend the easy interpretation and the wide coverage that Chinn and Ito (2006) provide. The natural alternative is to use actual capital flows as *de facto* measures of financial openness. However, actual flows are quite volatile from period to period, arguably too volatile to be accurately representing the generally slower moving changes in the underlying policies that are of interest to us here.

[15]To see that a stable policy does not need a stable exchange rate, consider Canada, which has had a floating exchange rate and open capital market for more than two decades. Throughout this period, its exchange rate has fluctuated, but its *policy* of floating exchange rates and open capital markets has remained the same.

[16]While we examine the full sample of countries, we note that rich economies, middle-income economies, and poor economies differ from one another in many ways that are neither well measured nor well understood. So, imposing constancy may entail questionable restrictions (even when unconditional distributions look broadly similar). Separating the income groups is the simplest way to allow them to differ. The income groupings are available at www.worldbank.org.

As we will see with more detail below, this reflects the fact that distributions are skewed to the left and bounded from below by zero. The middle income countries have both the highest means and the highest median.[17] The norm's maximum values are given in the next column. The largest value, 0.94, belongs to a middle income country: Mexico, which in 1976 abruptly ended its peso fix in exchange for greater monetary sovereignty. The standard deviations show that the norms vary widely throughout all of the subsamples.

The table also provides measures of skewness and kurtosis. As can be seen in the labeled columns, the norms, regardless of how they are split, appear to be strongly leptokurtotic and positively skewed. This can be seen more clearly in Figure 3. The top chart plots non-parametric kernel density estimates of the distributions of the norm for each of the three income groups, and the bottom chart provides comparable plots for each decade in the sample. Both charts show plainly that there are many observations where trilemma policy changes are either small or zero. While very small values are slightly more predominant in the high-income economies, they are prevalent across all three income groups. Likewise, we see a greater concentration of small values in the eighties and the naughts that in the other decades, but the skewness and leptokurtosis are striking in all decades.

[INSERT FIGURE 3 ABOUT HERE]

Figure 4 graphs the means of the norms over time. The top four charts plot the norms for each of the income groups; and, the red lines show the mean adjusted norms constructed using the interest rate correlation measure of

[17] However, Hodges and Lehmann (1963) estimates of the median differences are all zero.

sovereignty, while the blue lines use the new, implied trilemma measure of sovereignty.[18]

[INSERT FIGURE 4 ABOUT HERE]

Overall, the trilemma policies appear to be more stable when that stability is assessed using the new, trilemma-implied monetary sovereignty measure. However, both measures allow us to see the rise in policy changes in low and middle income countries in the nineties – around the time of the Asian financial crisis. Likewise, both measures clearly indicate the policy instability occurring in high-income countries after the fall of Bretton Woods. Throughout most of the remainder of the paper we calculate the norms using the new, trilemma-implied measure of monetary sovereignty.[19]

3.2. Archetypes

Next, we explore how the norms differ across the types of trilemma arrangements. We assign observations to four different types of arrangements based on their semblance to one of four "archetypes:" a '*Hong-Kong*' type,

[18] Appendix Table A1 splits the sample according to the dates of some of the key crises that occur during the period: the Mexican Crisis (1994), the Southeast Asian Crisis (1997), and the Argentine Crisis (2002). Summary statistics are provided for each of the subsamples. For the high-income countries, the table reports lower means, medians, maxima, and standard deviations in the later part of the sample than in the early part, regardless of where the split is made. However, the estimated Hodges-Lehmann differences in medians again all equal zero, and the differences for the other income groups are less uniform.

[19] At times, it is the very large changes in policy that are of most interest. So, we separately examine the incidence of large observations. Appendix Table A2 provides data on the largest decile of adjusted norms. The table lists the number of these large observations in each year, by income group and for the full sample. In each cell within the table, the numerator gives the number of the large observations, while the denominator gives the total number of observations. Overall, the pattern of large policy changes follows the pattern of the means. The richest economies have the fewest large changes in their trilemma policies, while the middle-income group has the highest proportion of large changes.

with exchange rate stability and open capital markets; a '*China*' type, with exchange rate stability and monetary sovereignty; a '*U.S.*' type with open financial markets and monetary sovereignty; and a '*Middle*' type, with a modest degree of all three characteristics.

We use the simple geometry of the trilemma to describe the types of arrangements more precisely. Letting j= '*Hong Kong*', '*China*', '*U.S.*', '*Middle*', we define $type_j$ such that R_j takes on the values: $(1, 1, 0)$, $(1, 0, 1)$, $(0, 1, 1)$, and $(\frac{2}{3}, \frac{2}{3}, \frac{2}{3})$. Each of these four values of R_j represents a point on the frontier of the feasible set defined by the trilemma. The first three points represent the three corners corresponding to the '*Hong Kong*,' '*China*,' and '*U.S.*' archetypes described above, and the last point represents the '*Middle*' of the feasible frontier. Then, we define country i's type in period t by its proximity to one of the four points. Specifically, we let:

$$j = \underset{j}{argmin} ||(R_{i,t} - R_j)||$$

$$type_{i,t} \overset{def}{=} type_j.$$

That is, the observation's type is defined by the one that minimizes the distance between the observation and the archetype.[20]

Throughout much of the modern period, the most common arrangement in this taxonomy is the '*China*' type, with its relatively stable exchange rates and a relatively high degree of monetary sovereignty. The second most common arrangement type is the '*Middle*.' The number of '*Middle*'

[20]Using this definition of assigned types, Figure A2 in the appendix shows the number of economies in each year of each type. In our sample, the observations of the Chinese, Hong Kong, and U.S. economies do not precisely mimic the zero or one values of their corresponding archetypes, but they are close.

observations rose through the early nineties as many '*China*' type economies began to relax some of their capital controls. The number of economies of the '*Hong Kong*' type has been rising fairly steadily since the nineties. The number of economies of the '*U.S.*' type has risen throughout the period, though less steadily.[21]

[INSERT FIGURE A2 ABOUT HERE]

Next, we examine the stability of the archetypes by looking at the norms in each category. Specifically, for each observation, we note the archetype and observe the extent of the trilemma policy change over the subsequent year.

The third panel of Table 1 summarizes the adjusted norm for the four types of arrangements. As shown, the economies that fall within the fixed exchange rate archetypes, '*China*' and '*Hong Kong*', are the ones that have the smallest means and medians.[22] Notably, the median of the observations in the '*Hong Kong*' archetype is zero. Underlying this statistic is the fact that about two-thirds of the norms in the *Hong Kong*' category are zero.[23]

[21]These findings can be interpreted as confirmation that there has been no sustained 'hollowing out of the middle,' where the 'middle' is now defined in the three-dimensional context of the trilemma. Suggested first in the nineties, the 'hollowing out' argument was that increasing capital mobility would make intermediate exchange rate regimes unsustainable; so governments would be forced to choose between zero and full exchange rate stability. Frankel et al. (2001) and others later refuted the argument empirically by noting that policies of modest exchange rate stability were holding their own against the extremes of fixity and floating. The bulk of the literature focused exclusively on the single dimension of exchange rate policy. Here, one can define the 'middle' and the 'poles' in terms of all three policy dimensions we likewise find that the hollowing out idea is not supported. The approach builds on the findings of Aizenman, Chin, and Ito, who show that emerging market economies have moved toward a blend of policies.

[22]Hodges-Lehmann estimates (not reported in the table) of the pseudo median differences between each archetype and the remainder are nonzero for all four categories.

[23]Note that there is nothing inherent in the 'Hong Kong' type that necessitates that it is the most stable policy configuration. We can see this by way of example. Consider Argentina in 2001, when it was characterized as a 'Hong Kong' type in 2001. This archetype's policy triad was not sustained. Although Argentina retained a fixed exchange

That is, economies with relatively fixed exchange rates and open capital markets often keep their policies the same from one year to the next. Correspondingly, '*Hong Kong*' is also the archetype with the greatest leftward skewness, shown in the designated column. The '*China*' archetype, which has relatively closed financial markets, is also heavily skewed to the left, with a low median, and many (about forty percent) of its norms equal to zero. The '*U.S.*' archetype, in which exchange rates are flexible and financial markets are open, and the '*Middle*' archetype, which has some of that openness and flexibility, have higher means and medians. That is, not only do these last two archetypes have more variable exchange rates, they also have more variable trilemma policies.

The bottom rows of Figure 4 illustrate how stability has changed over the modern period for each of the archetypes. Despite the obvious peaks in the mean adjusted norms of the '*China*' and '*Hong Kong*' archetypes in the late nineties, these archetypes (which have exchange rate stability in common) exhibit the smallest overall policy changes; and, their relative stability has been largely sustained throughout the global financial crisis. While the norms of the '*U.S.*' archetype countries have fallen over the modern era as a whole, they – along with the norms of the *'middle'* category – have been relatively high.

4. Panel Regressions

This section uses panel estimates to explore the relationship between stability and the underlying trilemma policies. The bounded nature of the

rate between 2001 and 2002, it changed its financial openness and monetary sovereignty considerably, which gave it a large norm: $n_{i,t} = 0.59$. This value differs only slightly from Indonesia's large norm at the time of the Asian Crisis.

adjusted norm raises a number of econometric issues that render the use of linear models potentially problematic.[24] Papke and Wooldridge (2008) propose a solution to this problem in a panel context, and the estimation here relies on their approach. Their solution employs a generalized estimation equation (GEE) in a balanced panel.

Using this approach, two specifications are estimated with a balanced panel of 96 countries between 1985–2010.[25] Both specifications relate the norm to the underlying trilemma policies and to official holdings of foreign exchange reserves.[26]

The first specification relates the adjusted norm to past reserves and past measures of exchange rate stability and of financial openness. The second specification also includes lagged reserves, but instead of including the measures of exchange rate stability and openness, it includes dummies for the economy's lagged archetype.

Specifically, GEE estimates are provided for two versions of Papke and Wooldridge's (2008) fractional panel model:

$$E(n_{it}|\mathbf{x}_{i1},\ldots,\mathbf{x}_{iT}) = \Phi(\kappa_t + \mathbf{x}_{it}\beta + \bar{\mathbf{x}}_i\lambda)$$

where $\mathbf{x}_{it}$ is the vector of explanatory variables; $\bar{\mathbf{x}}_i$ is the corresponding vector of country-specific means; κ_t, β, and λ are scaled coefficients; and, the time subscript in κ_t indicates the use of a complete set of time dummies. The

[24] For details see Papke and Wooldridge (1996).

[25] An appendix lists the countries that are included in the balanced panel.

[26] The consideration of reserves reflects a long tradition of studying their links to trilemma policies. Beginning with the early work on optimal reserves in a stochastic setting (for example: Kenen and Yudin (1965) and Heller (1966)), economists have modeled reserves as potentially reducing the probability or cost of devaluations, of speculative attacks, and of sudden stops. Their inclusion here allows for such a role. Data are taken from the World Bank's World Development Indicators.

inclusion of $\bar{\mathbf{x}}_i$ allows for time-constant, unobserved country effects that may be related to our other regressors, while it avoids the incidental parameters problem raised by cross-sectional dummies in this context.[27] In the first specification, $\mathbf{x}_{it} = [\rho_{i,t-1}, s_{i,t-1}, f_{i,t-1}]$; and, in the second specification, $\mathbf{x}_{it} = [\rho_{i,t-1}, D_{\text{`}China\text{'},i,t-1}, D_{\text{`}HongKong\text{'},i,t-1}, D_{\text{`}U.S.\text{'},i,t-1}]$, where ρ is the ratio of official reserves to GDP, and D_j indicates a dummy variable for $type_{i,t} = R_j$.[28]

Panel estimates are first presented using the full range of policy changes, and the three income groups are treated separately. The focus then turns to large policy changes exclusively, where 'large' is defined in terms of several cutoffs of the value of the norm.

4.1. Estimation by Income Group

The top two panels of Table 2 provide the estimation results from the two specifications for each income group. The top panel gives the estimates from the specification using the exchange rate stability and financial market openness as regressors; and, the second panel gives the estimates from the specification that makes explicit use of the archetypes. Each pair of columns gives the estimated coefficients, along with their standard errors – which are robust to second order misspecification – and the partial effects, averaged across the population (APEs), with bootstrapped standard errors.[29]

[INSERT TABLE 2 ABOUT HERE]

The results for low-income economies are given in the first pair of columns; and, the first row in each panel gives the estimates for reserves as a fraction

[27] Lancaster (2000) provides a survey of the literature on the incidental parameters problem.

[28] Note that $R_{\text{`}Middle\text{'}}$ is subsumed by the constants in the second specification.

[29] For a discussion of APEs, see chapter 2 in Wooldridge (2010).

of GDP. In both specifications, the estimated coefficients are negative and statistically significant at the five percent level. The estimated APEs, which (unlike the raw coefficients) can be compared across specifications, are of roughly similar magnitudes: 0.34 in the first specification and 0.40 in the second. These estimates imply that in low-income economies greater reserves tend to come with greater trilemma policy stability.

The low-income estimates for the first specification's remaining variables, the degree of exchange rate stability and the degree of financial openness, are given with their standard errors in the subsequent rows of the top panel. None is statistically significantly different from zero at any standard confidence level. The low-income archetype estimates are given in the remaining rows of the second panel. As shown, the coefficient on the 'Hong Kong' archetype is positive and mildly statistically significant. This implies that (conditional on reserves), the combination of open capital markets and fixed exchange rates does not represent a particularly stable policy configuration among low income economies.

The next pair of columns provides the estimates for the middle income economies. Here, reserves are no longer statistically significant. However, in the first specification, we see that (in the third row) exchange rate stability has a negative coefficient and is mildly significant. That is, conditional on the reserves and the degree of financial openness, we are somewhat more likely to find smaller policy disruptions in middle-income economies when they have relatively more stable exchange rates. In the second specification, we see that the middle-income coefficient on the 'USA' archetype is positive and and mildly significant. That is, conditional on reserves and financial openness, larger policy changes are found here when exchange rates are flexible. The positive coefficient (implying a higher norm) on the 'U.S.'

archetype in this specification goes hand in hand with the negative coefficient (implying a lower norm) on exchange rate stability in the first specification.

The third pair of columns gives the estimates for the high-income economies. The estimated coefficients on reserves are positive and significant in both specifications. In the first specification, the estimated coefficient on exchange rate stability is negative, as it is in for middle-income economies, and here it is statistically significant at the one percent level. Correspondingly, in the second specification, the coefficient on the 'Hong Kong' archetype is negative and statistically significant at all confidence levels.

Finally, the last pair of columns gives the estimates for the full sample. Taken as a whole, reserves lose all significance in the first specification, but the estimated coefficient on exchange rate stability is negative and significant. Likewise, the coefficient on reserves is small and insignificant in the second specification, but the coefficient on the 'Hong Kong' archetype is again negative and strongly statistically significant.

4.2. Large Norms

The bottom of the table provides the results from the same specifications estimated only for 'large' policy changes. The first pair of columns report estimates using norms from the top decile of the distribution; and, subsequent columns provide estimates where the definition of 'large' is broadened to include additional deciles, until all the values of the norm above the median are included.

The first row of each panel again gives the estimated reserve parameters. In both specifications, and for all definitions of 'large', the estimated coefficients on reserves are negative, though there is only mild statistical significance for the top decile estimates, and none elsewhere. Recall that the

APE estimates, unlike the raw coefficient estimates, can be compared across specifications. In the first specification the estimated APE is -0.37, and in the second specification it is -0.42. These values do not differ markedly from the earlier low-income APE estimates of -0.34 and -0.40.

The negative link between reserves and trilemma stability that we see here in times of instability, and for low-income economies, may reflect a greater incidence of limitations on governments' access to international financial markets. With limited access to credit, the governments in such economies must rely more heavily on their own reserves when funds are needed to smooth policies.[30]

The estimates for exchange rate stability and financial market openness are given next. The estimated exchange rate stability coefficients are uniformly positive for the large norms. That is, there is some tendency to find large policy changes in conjunction with greater exchange rate stability. These positive estimates contrast with the negative full sample estimates above. The statistical significance of the exchange rate is limited here to the samples that include the top 30 percent and the top 40 percent. One possible interpretation of these results is that fixed exchange rates are usually part of relatively stable policies, but when they are associated with policy changes, those changes are somewhat large.

The estimated coefficients on financial market openness are given next. The coefficients again are all negative; and they are significant here at the one percent level for all but the top decile. This tells us that, conditional on reserves and the degree of exchange rate stability, trilemma policies are

[30] As mentioned above, such policy smoothing is typically optimal in models with convex policy costs. See, for example, Pina (2012) for a model of developing country reserves in a monetary policy context.

more stable when financial markets are open.

Among the archetypes, given next, only the *'China'* estimates are statistically significant at standard confidence levels. The *'China'* estimates are uniformly positive, and they are statistically significant for all definitions of 'large', except the top decile. The positive *'China'* coefficients echo the earlier findings for exchange rate stability and financial openness in that the *'China'* archetype represents the combination of relatively stable exchange rates and relatively closed financial markets. That is, both these qualities are associated with relatively large policy changes.

5. Conclusions

Underlying this paper is a willingness to use the constraint of the classic, open-economy trilemma and to draw out some of its implications for empirical work on the stability of trilemma policies. The simple geometry of the trilemma is used to provide a univariate gauge of the stability of a country's multidimensional international macroeconomic policies. The new gauge is bounded by the constraints of the trilemma itself, and it is non-Gaussian. Most importantly, the distribution is asymmetric. Future studies of trilemma policy stability – whether studies of its determinants or its consequences – should recognize and incorporate this fundamental asymmetry.

In addition to the new trilemma stability gauge, the paper provides a new, implicit measure of monetary sovereignty; and it illustrates a framework for characterizing international macroeconomic arrangements in terms of their semblance to definitive policy archetypes. The monetary sovereignty measure is constructed from the trilemma's constraint in conjunction with existing measures of exchange rate stability and international financial open-

ness. The international macroeconomic policy characterizations stem from their positioning within the trilemma's policy space.

The paper's approach and its resulting measures are used here to characterize the international macroeconomic arrangements of the modern era. The measures indicate that international macroeconomic policies have been most stable in settings of relatively fixed exchange rates and open financial markets. Using the new monetary measures, it appears that for many countries monetary sovereignty has been both somewhat greater and somewhat less erratic than previously had been thought. Finally, when attention is restricted to large policy changes or to low-income economies, the stability of international macroeconomic policies also appears to be linked to official holdings of foreign exchange reserves.

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Table 1: Norms Across Income Groups, Over Time, and by Archetype

	Norm	**Mean**	**Median**	**Max.**	**Min.**	**St. Dev.**	**Skew.**	**Kurt.**	**Obs.**
Across Income Groups									
Low Income Economies	Imp.	0.10	0.04	0.76	0.00	0.14	1.96	7.18	1,307
	Corr.	0.13	0.10	0.72	0.00	0.12	1.93	8.20	1,093
Middle Income Economies	Imp.	0.11	0.06	0.94	0.00	0.15	2.05	7.53	2,821
	Corr.	0.13	0.11	0.76	0.00	0.13	1.88	7.61	2,399
High Income Economies	Imp.	0.09	0.05	0.77	0.00	0.12	2.10	8.78	1,570
	Corr.	0.11	0.09	0.67	0.00	0.09	1.54	7.30	1,436
Over Time									
1970s	Imp.	0.12	0.05	0.94	0.00	0.16	1.71	5.76	1,099
	Corr.	0.14	0.11	0.68	0.00	0.13	1.52	5.58	761
1980s	Imp.	0.09	0.04	0.93	0.00	0.14	2.68	11.34	1,359
	Corr.	0.11	0.09	0.76	0.00	0.12	2.56	11.69	1,148
1990s	Imp.	0.12	0.06	0.78	0.00	0.14	1.84	6.85	1,580
	Corr.	0.14	0.11	0.72	0.00	0.12	1.81	7.68	1,445
2000s	Imp.	0.08	0.04	0.88	0.00	0.12	2.18	8.94	1,697
	Corr.	0.12	0.10	0.71	0.00	0.10	1.71	8.03	1,601
By Archetype									
'China' Archetype	Imp.	0.09	0.02	0.75	0.00	0.14	2.16	8.05	2,365
'Hong Kong' Archetype	Imp.	0.06	0.00	0.77	0.00	0.13	2.89	11.79	790
'U.S. Archetype'	Imp.	0.12	0.08	0.74	0.00	0.13	1.80	6.84	2,015
'Mid Archetype'	Imp.	0.16	0.09	0.94	0.00	0.18	1.98	6.76	565
All Economies, Years & Archetypes									
All	Imp.	0.10	0.05	0.94	0.00	0.14	2.11	8.09	5,735
	Corr.	0.12	0.10	0.76	0.00	0.12	1.94	8.39	4,955

Note: The interest rate based correlation measure of trilemma stability (Corr.) is calculated using data from Aizenman, Chinn, and Ito (2010), who in turn use Shambaugh's (2004) interest rate correlation-based measure of monetary sovereignty. The implied norm (Imp.) is calculated using the new measure of monetary sovereignty described in section 2.2. The dataset consists of 177 countries and the maximum sample extends from 1971 to 2010 (the dataset is unbalanced). Income group classifications are from the World Bank (January 2011), available at www.worldbank.org.

Table 2: GEE Estimates

By Income Group

	Low Income		**Middle Income**		**High Income**		**All Countries**	
	Coeff.	APE	Coeff.	APE	Coeff.	APE	Coeff.	APE
Specification I								
International Reserves	−2.563**	− 0.339***	0.113	0.021	0.753**	0.095*	0.184	0.030
	(1.039)	(0.117)	(0.255)	(0.046)	(0.348)	(0.053)	(0.223)	(0.034)
Exchange Rate Stability	0.280	0.037	−0.315*	−0.060**	−1.085***	−0.137***	−0.379**	−0.063***
	(0.469)	(0.035)	(0.170)	(0.023)	(0.164)	(0.023)	(0.154)	(0.018)
Financial Openness	−0.143	−0.019	−0.175	−0.033	−0.047	−0.006	−0.202	−0.033**
	(0.131)	(0.019)	(0.172)	(0.021)	(0.168)	(0.025)	(0.131)	(0.015)
Specification II								
International Reserves	−2.712**	−0.401***	0.135	0.026	0.654*	0.077	0.140	0.023
	(1.046)	(0.116)	(0.275)	(0.046)	(0.349)	(0.052)	(0.215)	(0.033)
'China' Archetype	−0.011	−002	0.058	0.011	−0.329	−0.039*	0.003	0.001
	(0.068)	(0.012)	(0.059)	(0.01)	(0.261)	(0.02)	(0.052)	(0.008)
'Hong Kong' Archetype	0.325*	0.049	−0.243	−0.046**	−0.806***	−0.094***	−0.477***	−0.079***
	(0.179)	(0.034)	(0.185)	(0.023)	(0.150)	(0.017)	(0.129)	(0.014)
'USA' Archetype	−0.036	−0.05	0.172*	0.033**	0.026	0.003	0.095	0.016*
	(0.114)	(0.015)	(0.098)	(0.013)	(0.105)	(0.011)	(0.067)	(0.009)
Obs. (Countries)	2,400 (96)		450 (18)		1,300 (52)		650 (26)	

By Centile of Norms

	Top 10%		**Top 20%**		**Top 30%**		**Top 40%**		**Top 50%**	
	Coeff.	APE	Coeff.	APE	Coeff.	APE	Coeff.	APE	Coeff.	APE
Specification I										
International Reserves	−0.971	−0.37*	−0.406	−0.141	−0.359	−0.116	−0.395	−119*	−0.267	−0.072
	(0.617)	(0.219)	(0.329)	(0.108)	(0.299)	(0.079)	(0.310)	(0.070)	(0.291)	(0.063)
Exchange Rate Stability	0.047	0.018	0.131	0.045	0.177**	0.057**	0.242***	0.073***	0.054	0.015
	(0.077)	(0.037)	(0.081)	(0.034)	(0.087)	(0.028)	(0.091)	(0.026)	(0.099)	(0.025)
Financial Openness	−0.147	−0.056	−282***	−0.098***	−0.314***	−0.102***	−0.298***	−0.09***	−0.367***	−0.1***
	(0.119)	(0.051)	(0.108)	(0.031)	(0.105)	(0.026)	(0.110)	(0.023)	(0.119)	(0.03)
Specification II										
International Reserves	−1.118*	−0.424**	−0.371	−0.128	−0.228	−0.073	−0.282	−0.083	−0.197	−0.052
	(0.591)	(0.210)	(0.331)	(0.106)	(0.307)	(0.077)	(0.334)	(0.071)	(0.301)	(0.062)
'China' Archetype	0.036	0.014	0.113**	0.039**	0.119**	0.038***	0.121***	0.036***	0.131***	0.035***
	(0.059)	(0.026)	(0.052)	(0.017)	(0.047)	(0.014)	(0.046)	(0.012)	(0.049)	(0.011)
'Hong Kong' Archetype	−0.113	−0.043	−0.099	−0.034	−0.009	−0.003	0.069	0.02	−0.186	−0.049**
	(0.083)	(0.037)	(0.088)	(0.027)	(0.075)	(0.023)	(0.095)	(0.023)	(0.119)	(0.021)
'USA' Archetype	0.118*	0.045	0.053	0.018	0.072	0.023	0.064	0.019	0.053	0.014
	(0.062)	(0.03)	(0.055)	(0.021)	(0.051)	(0.017)	(0.055)	(0.014)	(0.057)	(0.013)
Obs. (C)	240 (68)		483 (89)		722 (92)		963 (92)		1208 (94)	

Note: Dependent variable is the norm implied by the trilemma. Time averages of the explanatory variables and time dummy variables are included in the regressions. Explanatory variables are lagged one period. Standard errors (in parentheses) are robust to conditional variance and serial correlation. APE standard errors are bootstrapped (500 runs). Single, double, and triple asterisks denote statistical significance at the ten, five, and one percent level.

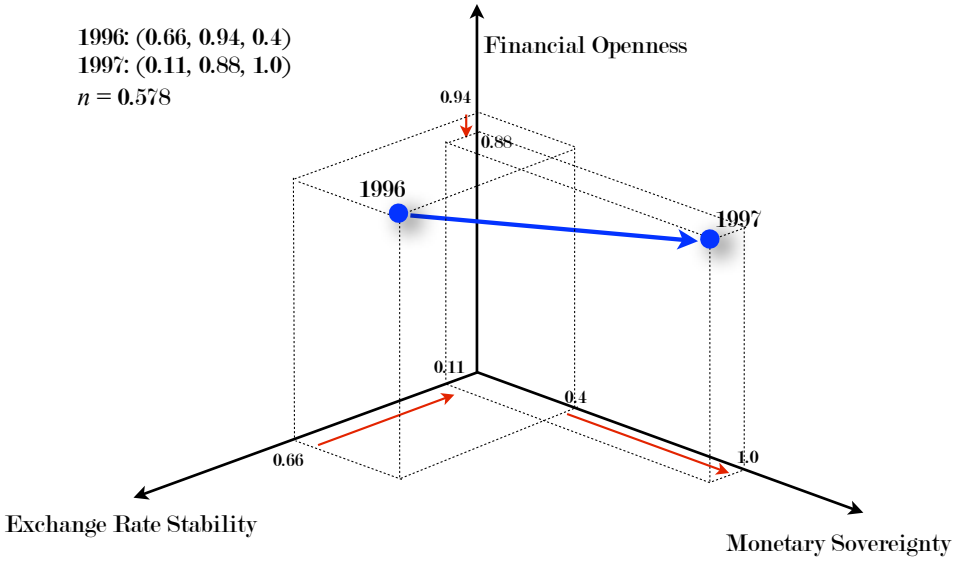

Figure 1: Indonesia 1996–97. The figure displays the two data points underlying a single observation of the adjusted norm, $n_{i,t}$. As is well-known, Indonesia experienced a substantial increase in its exchange rate variability and a small reduction in its financial openness during the crisis, while it increased its monetary sovereignty considerably. These changes are indicated by the vector shown between the observations for 1996 and for 1997. The normalized length of the vector measures the overall change in the policy triad.

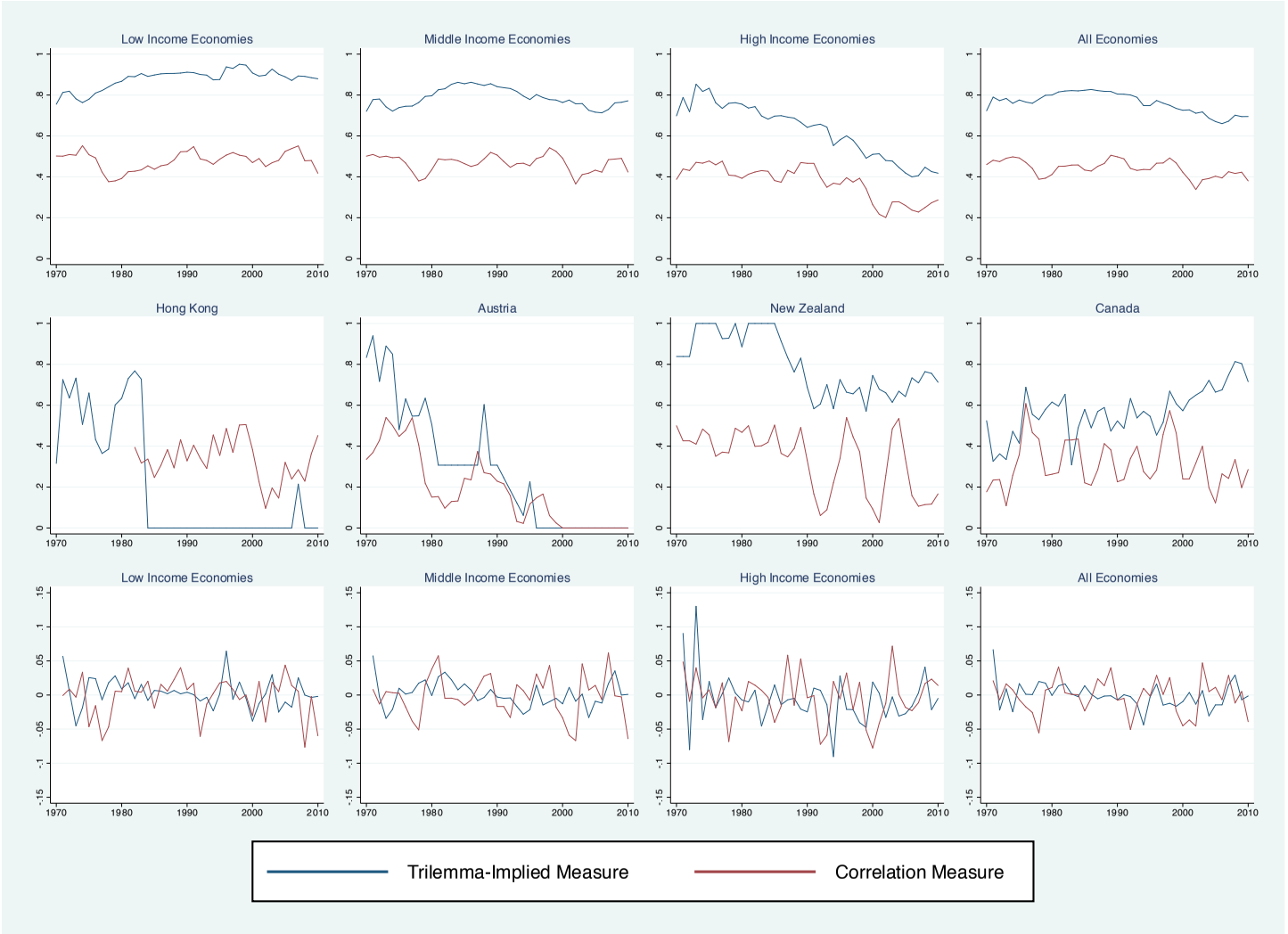

Figure 2: Monetary Sovereignty. The top row provides plots of the trilemma-implied and the correlation-based measures of monetary sovereignty for three income groups and for the full sample. Looking at the means, the new, implicit measure suggests a greater degree of monetary sovereignty than does the standard, interest rate measure. The middle row plots the two gauges for Hong Kong, Austria, New Zealand and Canada. These examples show that the two measures can differ in either direction. Finally, the bottom row provides plots of the *changes* in the two measures of monetary sovereignty. Using the new measure, it is now easy to see the monetary upheaval many countries (especially the high income ones, shown in the row's third plot) experienced in the wake of the Bretton Woods breakdown.

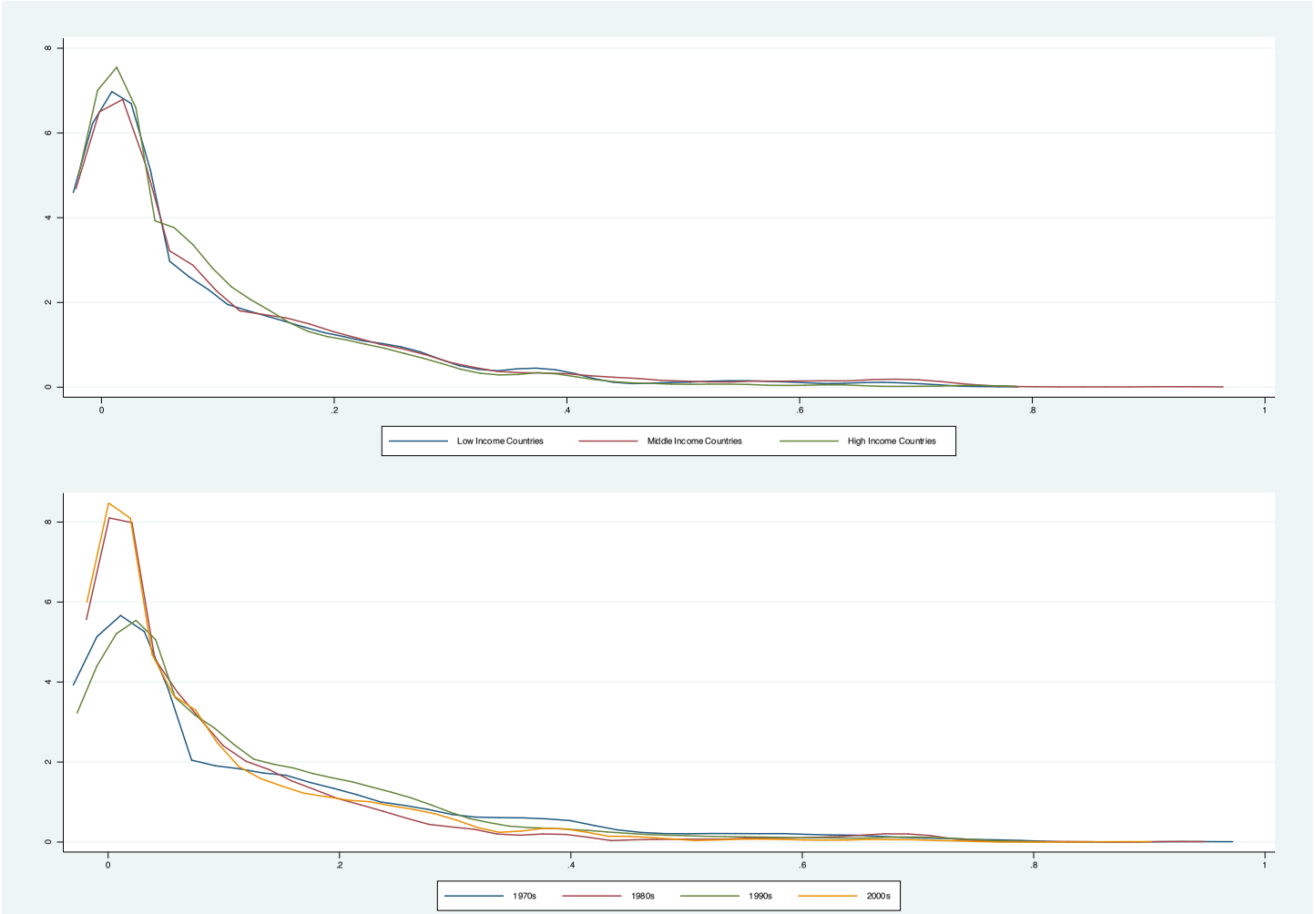

Figure 3: Norm Kernel Densities. The first panel of the figure provides plots of non-parametric kernel density estimates of the distributions of the norm for each of the three income groups, and the second panel provides comparable plots for each decade in the sample. Both plots show plainly that there are many observations where trilemma policy changes are either small or zero. While very small values are slightly more predominant in the high-income economies, they are prevalent across all three income groups. Likewise, we see a greater concentration of small values in the eighties and the naughts that in the other decades, but the skewness and leptokurtosis are striking in all decades.

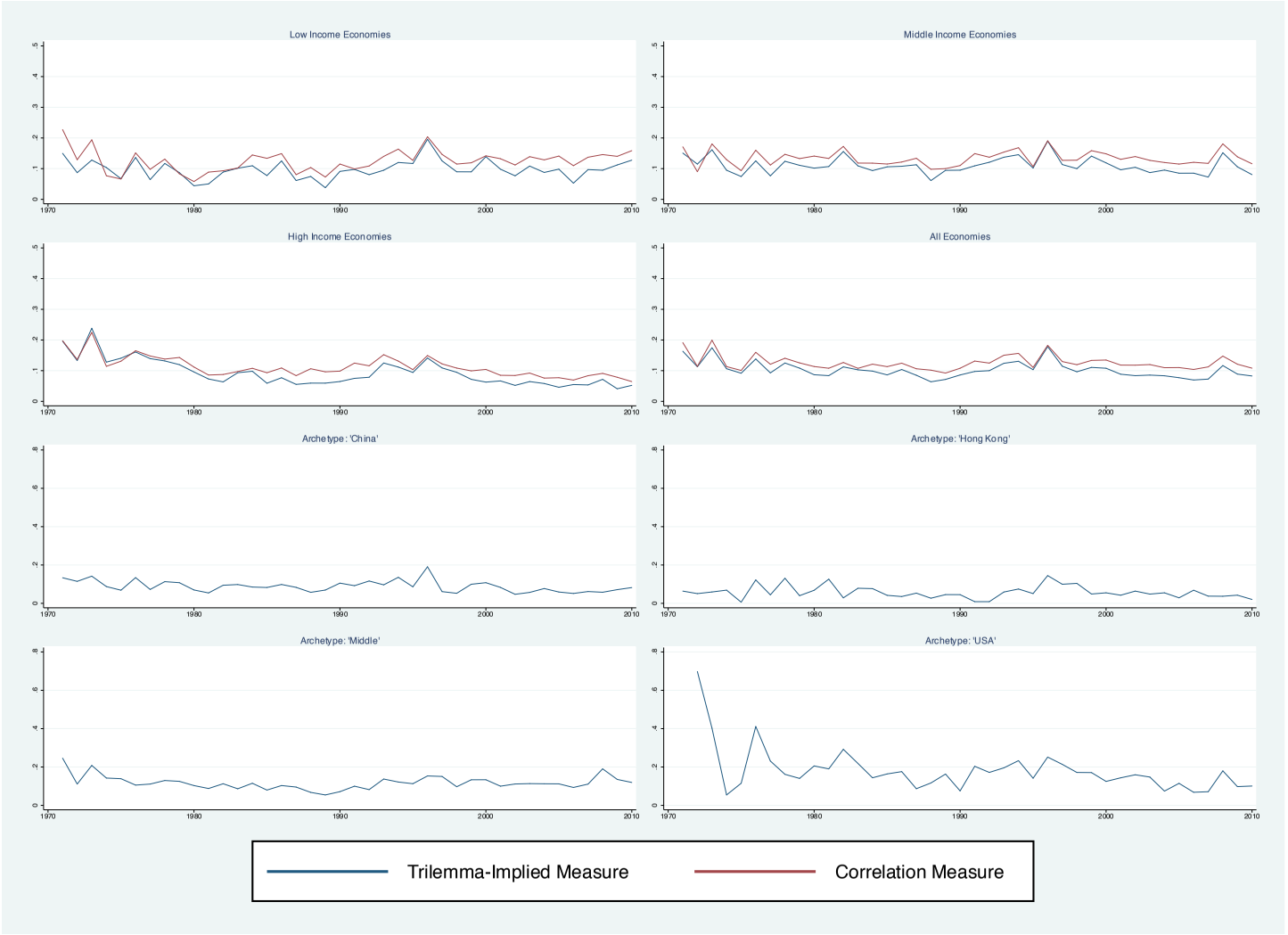

Figure 4: Mean Norms. The top four graphs show plots of the trilemma-implied and correlation-based norms for the income groups and the entire sample. Overall, the trilemma policies appear to be more stable when that stability is assessed using the former. The bottom four graphs show plots of the norms over time by international policy archetype. Despite the obvious peaks in the norms of the '*China*' and '*Hong Kong*' archetypes in the late nineties, these archetypes, which have exchange rate stability in common, exhibit the smallest overall policy changes; and, their relative stability has been largely sustained throughout the global financial crisis. While the norms of the '*U.S.*' archetype countries have fallen over the modern era as a whole, they – along with the norms of the *'middle'* category - have been relatively high.

A. ON-LINE APPENDIX

Table A1: Norms Before and After Recent Crises

	Group	Mean	Median	Max.	Min.	St. Dev.	Skew.	Kurt.	Obs.
1971–1994	LIC	0.09	0.02	0.69	0.00	0.14	2.18	8.00	689
	MIC	0.11	0.05	0.94	0.00	0.16	2.11	7.49	1390
	HIC	0.10	0.06	0.77	0.00	0.13	2.08	8.56	818
	All	0.10	0.05	0.94	0.00	0.15	2.17	8.12	2897
1995–2010	LIC	0.11	0.06	0.76	0.00	0.13	1.72	6.32	618
	MIC	0.11	0.06	0.88	0.00	0.14	1.93	7.23	1431
	HIC	0.07	0.03	0.68	0.00	0.10	1.97	7.38	789
	All	0.10	0.05	0.88	0.00	0.13	1.98	7.63	2838
1971–1997	LIC	0.10	0.03	0.69	0.00	0.14	2.01	7.15	790
	MIC	0.12	0.05	0.94	0.00	0.16	2.00	6.99	1617
	HIC	0.11	0.06	0.77	0.00	0.13	2.00	8.32	938
	ALL	0.11	0.05	0.94	0.00	0.15	2.05	7.61	3345
1998–2010	LIC	0.10	0.05	0.76	0.00	0.13	1.86	7.10	517
	MIC	0.10	0.06	0.88	0.00	0.14	2.06	8.02	1204
	HIC	0.06	0.02	0.68	0.00	0.10	2.23	8.60	669
	All	0.09	0.05	0.88	0.00	0.13	2.12	8.47	2390
1971–2002	LIC	0.10	0.04	0.76	0.00	0.14	1.97	7.15	975
	MIC	0.11	0.06	0.94	0.00	0.16	2.00	7.07	2057
	HIC	0.10	0.06	0.77	0.00	0.12	1.98	8.09	1178
	All	0.11	0.06	0.94	0.00	0.15	2.05	7.65	4210
2003–2010	LIC	0.09	0.05	0.72	0.00	0.12	1.86	6.93	332
	MIC	0.10	0.05	0.88	0.00	0.13	2.09	8.50	764
	HIC	0.05	0.01	0.68	0.00	0.09	2.56	11.48	429
	All	0.08	0.04	0.88	0.00	0.12	2.19	9.07	1525

Note: The table splits the sample according to the dates of some of the key crises that occur during the period: the Mexican Crisis (1994), the Southeast Asian Crisis (1997), and the Argentine Crisis (2002). Summary statistics are provided for each of the subsamples. For the high-income countries, the table reports lower means, medians, maxima, and standard deviations in the later part of the sample than in the early part, regardless of where the split is made.

Table A2: Norm Values in the Last Decile by Income Group

Year	Low Income	Middle Income	High Income	All
1971	9/26	12/48	9/29	30/103
1972	3/27	8/48	3/29	14104
1973	8/28	12/50	13/30	33/107
1974	4/27	4/51	3/31	11/110
1975	2/27	4/49	5/31	11/107
1976	4/26	7/51	4/31	15/108
1977	1/26	2/52	6/34	9/112
1978	4/26	7/53	4/36	15/115
1979	1/26	8/53	4/36	13/115
1980	1/27	7/55	2/36	10/118
1981	1/27	7/56	1/36	9/119
1982	3/30	9/58	1/37	13/125
1983	3/31	9/63	3/37	15/131
1984	3/31	4/65	2/37	9/133
1985	1/33	8/66	1/37	10/136
1986	4/33	7/70	3/37	14/140
1987	1/33	6/70	0/39	7/142
1988	1/33	2/70	2/39	5/142
1989	3/35	8/71	1/39	8/145
1990	2/35	5/72	1/39	9/146
1991	2/34	5/73	2/39	9/146
1992	2/34	10/73	1/39	13/146
1993	1/34	12/73	5/40	18/147
1994	5/34	13/74	3/40	21/148
1995	6/33	8/75	0/40	14/148
1996	6/34	14/78	2/40	22/152
1997	4/35	15/88	4/48	23/171
1998	2/37	10/88	6/48	18/173
1999	3/38	15/88	4/48	22/174
2000	8/38	12/89	0/48	20/175
2001	3/37	8/87	5/48	16/172
2002	4/37	6/86	1/48	11/171
2003	4/38	8/87	2/48	14/173
2004	2/38	9/87	2/48	13/173
2005	5/37	6/87	1/47	12/172
2006	1/37	5/84	3/47	9/168
2007	4/36	4/84	2/48	10/167
2008	1/36	17/83	3/48	21/167
2009	6/37	10/83	1/48	17/168
2010	5/36	3/83	1/47	9/166
Total	131/1,307	326/2,821	116/1,607	573/5,735
(%)	10.0	11.5	7.2	10.0

Note: Norm values over 0.2753 are in the last decile in the sample. Each numerator gives the number of such norms in the relevant portion of the sample, while the denominator gives the corresponding number of countries. Overall, the pattern of large policy changes follows the pattern of the means. The richest economies have the fewest large changes in their trilemma policies, while the middle-income group has the highest proportion of large changes.

Table A3: Descriptive Statistics of Trilemma-Implied Norms by Country

	Mean	**Median**	**Max.**	**Min.**	**St. Dev.**	**Skew.**	**Kurt.**
Algeria	0.09	0.08	0.24	0.01	0.06	0.71	3.21
Argentina	0.18	0.11	0.68	0.00	0.21	1.53	3.96
Austria	0.05	0.00	0.30	0.00	0.09	1.84	4.87
Bahamas, The	0.01	0.00	0.24	0.00	0.05	4.80	24.04
Bangladesh	0.17	0.19	0.42	0.00	0.14	0.23	1.92
Barbados	0.02	0.00	0.24	0.00	0.07	3.18	11.08
Belgium	0.07	0.02	0.54	0.00	0.13	2.44	8.24
Benin	0.01	0.00	0.24	0.00	0.05	4.80	24.04
Bolivia	0.14	0.06	0.67	0.00	0.16	1.80	5.95
Botswana	0.15	0.16	0.32	0.01	0.09	0.12	1.85
Brazil	0.12	0.08	0.55	0.00	0.12	2.08	7.58
Burundi	0.10	0.08	0.34	0.01	0.08	1.50	4.78
Cameroon	0.03	0.00	0.24	0.00	0.08	2.41	6.80
Canada	0.06	0.06	0.15	0.01	0.04	0.65	2.85
Central African	0.02	0.00	0.24	0.00	0.07	3.18	11.08
Chad	0.02	0.00	0.24	0.00	0.07	3.18	11.08
Chile	0.13	0.07	0.43	0.00	0.13	1.32	3.42
China	0.12	0.01	0.66	0.00	0.19	1.73	5.06
Colombia	0.12	0.10	0.55	0.02	0.11	2.24	9.22
Comoros	0.00	0.00	0.06	0.00	0.01	4.80	24.04
Congo, Rep.	0.02	0.00	0.24	0.00	0.07	3.18	11.08
Costa Rica	0.12	0.09	0.32	0.00	0.09	0.63	2.26
Côte d'Ivoire	0.03	0.00	0.24	0.00	0.08	2.41	6.80
Denmark	0.11	0.07	0.31	0.00	0.12	0.57	1.74
Dominica	0.06	0.00	0.29	0.00	0.11	1.31	2.76
Dominican Republic	0.21	0.17	0.69	0.02	0.18	1.68	4.95
Egypt, Arab Rep.	0.25	0.13	0.88	0.00	0.28	0.73	2.26
El Salvador	0.11	0.00	0.64	0.00	0.18	1.73	4.77
Equatorial Guinea	0.02	0.00	0.24	0.00	0.07	3.18	11.08

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	Mean	Median	Max.	Min.	St. Dev.	Skew.	Kurt.
Fiji	0.09	0.07	0.24	0.01	0.07	0.77	2.32
Finland	0.07	0.01	0.38	0.00	0.10	1.78	5.83
France	0.07	0.02	0.30	0.00	0.09	0.98	2.87
Gabon	0.01	0.00	0.24	0.00	0.05	4.80	24.04
Gambia, The	0.10	0.07	0.68	0.00	0.14	3.11	13.19
Germany	0.07	0.04	0.23	0.00	0.06	1.61	4.63
Greece	0.09	0.05	0.36	0.00	0.11	1.22	3.34
Grenada	0.04	0.00	0.41	0.00	0.11	2.23	6.66
Guatemala	0.13	0.11	0.65	0.00	0.14	2.20	9.05
Guyana	0.21	0.16	0.71	0.00	0.22	0.91	2.80
Honduras	0.15	0.07	0.69	0.00	0.21	1.64	4.58
Iceland	0.11	0.09	0.39	0.00	0.09	1.47	5.08
India	0.10	0.06	0.61	0.00	0.15	2.35	7.97
Indonesia	0.14	0.06	0.64	0.01	0.19	1.99	5.50
Ireland	0.08	0.03	0.47	0.00	0.12	1.89	5.82
Israel	0.11	0.09	0.40	0.00	0.09	1.57	5.83
Italy	0.10	0.07	0.44	0.00	0.13	1.02	3.11
Jamaica	0.18	0.17	0.43	0.00	0.12	0.30	2.12
Japan	0.06	0.05	0.17	0.00	0.05	0.70	2.37
Jordan	0.07	0.04	0.40	0.00	0.10	1.77	5.49
Kenya	0.10	0.05	0.50	0.00	0.12	2.00	6.98
Korea, Rep.	0.08	0.05	0.30	0.00	0.08	1.10	3.34
Madagascar	0.10	0.07	0.26	0.00	0.08	0.61	2.11
Malaysia	0.14	0.08	0.72	0.00	0.16	1.97	7.17
Maldives	0.11	0.06	0.47	0.00	0.12	1.36	4.26
Mali	0.01	0.00	0.24	0.00	0.05	4.80	24.04
Mauritania	0.12	0.12	0.39	0.00	0.09	0.80	3.84
Mauritius	0.11	0.08	0.28	0.00	0.08	0.76	2.41
Mexico	0.18	0.15	0.58	0.00	0.14	1.12	3.91

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Table A3 – continued from previous page

	Mean	Median	Max.	Min.	St. Dev.	Skew.	Kurt.
Morocco	0.07	0.05	0.21	0.00	0.07	0.82	2.24
Netherlands	0.09	0.00	0.50	0.00	0.14	1.66	4.93
Nicaragua	0.15	0.00	0.70	0.00	0.24	1.51	3.59
Niger	0.01	0.00	0.24	0.00	0.05	4.07	19.04
Nigeria	0.15	0.11	0.72	0.00	0.17	2.18	7.48
Norway	0.09	0.09	0.29	0.00	0.07	0.99	3.86
Pakistan	0.15	0.15	0.38	0.00	0.12	0.34	2.06
Panama	0.00	0.00	0.00	0.00	0.00	.	.
Papua New Guinea	0.11	0.09	0.30	0.00	0.09	0.77	2.53
Paraguay	0.16	0.11	0.68	0.00	0.17	2.22	7.27
Peru	0.17	0.14	0.54	0.00	0.14	1.09	3.54
Philippines	0.16	0.14	0.58	0.00	0.15	1.14	3.89
Portugal	0.08	0.02	0.50	0.00	0.13	1.71	5.26
Rwanda	0.10	0.06	0.31	0.00	0.09	0.97	2.98
Samoa	0.08	0.07	0.29	0.01	0.07	1.27	4.88
Saudi Arabia	0.02	0.00	0.31	0.00	0.06	3.74	17.26
Senegal	0.01	0.00	0.24	0.00	0.05	4.80	24.04
Seychelles	0.13	0.08	0.61	0.00	0.17	1.74	5.21
Sierra Leone	0.14	0.10	0.69	0.02	0.15	2.19	7.99
Singapore	0.11	0.07	0.54	0.00	0.13	1.91	6.49
Solomon Islands	0.14	0.08	0.65	0.00	0.17	1.50	4.64
South Africa	0.08	0.08	0.26	0.00	0.06	0.82	3.42
Spain	0.08	0.02	0.40	0.00	0.11	1.33	3.95
Sri Lanka	0.16	0.10	0.45	0.01	0.14	0.72	2.19
St. Lucia	0.05	0.00	0.29	0.00	0.11	1.59	3.57
St. Vincent and	0.02	0.00	0.24	0.00	0.07	3.18	11.08
Swaziland	0.01	0.00	0.24	0.00	0.05	4.80	24.04
Sweden	0.08	0.08	0.18	0.01	0.04	0.30	2.85
Thailand	0.10	0.05	0.52	0.00	0.13	2.08	6.69

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Table A3 – continued from previous page

	Mean	**Median**	**Max.**	**Min.**	**St. Dev.**	**Skew.**	**Kurt.**
Togo	0.00	0.00	0.00	0.00	0.00	.	.
Trinidad and Tobago	0.13	0.01	0.73	0.00	0.19	1.49	4.81
Tunisia	0.10	0.09	0.24	0.01	0.06	0.87	2.65
Turkey	0.09	0.07	0.25	0.01	0.07	0.68	2.18
Uganda	0.15	0.07	0.69	0.00	0.18	1.83	5.89
United Kingdom	0.09	0.05	0.27	0.00	0.08	1.01	2.85
Uruguay	0.16	0.16	0.45	0.01	0.11	0.77	3.18
Venezuela, RB	0.18	0.07	0.72	0.00	0.23	1.54	4.04
Zambia	0.13	0.11	0.53	0.01	0.11	2.05	8.27
Total	0.10	0.05	0.88	0.00	0.13	2.20	8.83

Note: The table reports descriptive statistics of the trilemma implied norm for each country in the dataset.

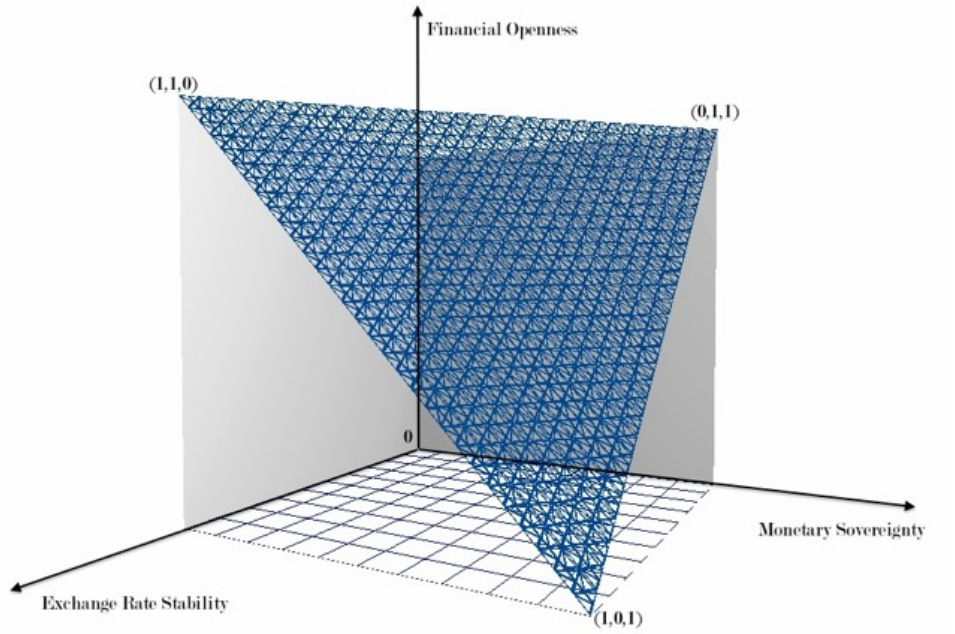

Figure A1: The Trilemma Constraint. This figure depicts a particular, linear form of the trilemma constraint; namely, with the indicators of exchange rate stability, financial openness, and monetary sovereignty each normalized between zero and one, they sum to two.

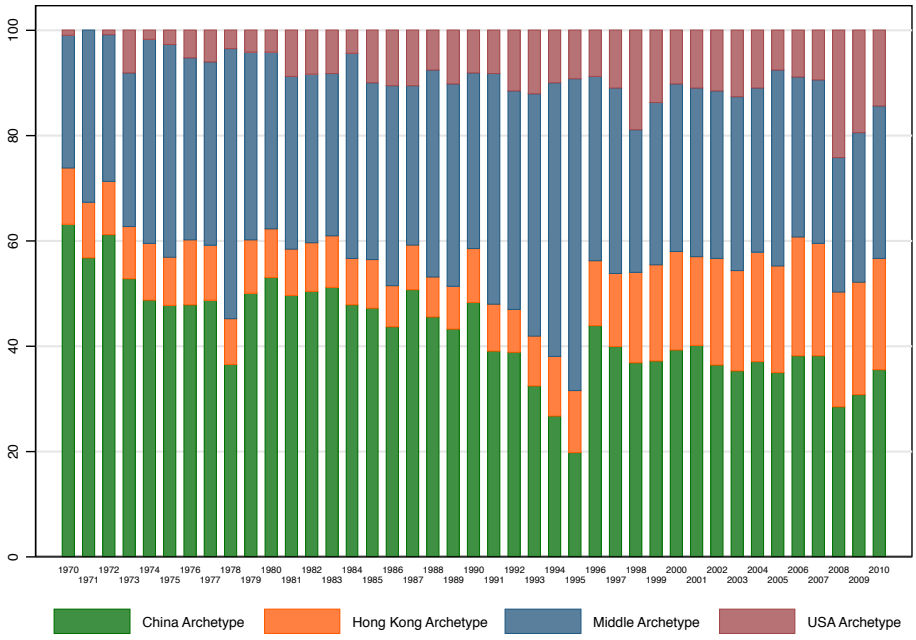

Figure A2: Countries per Archetype (percent). The figure shows the number of economies in each year of each type. Throughout much of the modern period, the most common arrangement in this taxonomy is the '*China*' type, with its relatively stable exchange rates and a relatively high degree of monetary sovereignty. The second most common arrangement type is the '*Middle*.' The number of '*Middle*' observations rose through the early nineties as many '*China*' type economies began to relax some of their capital controls. The number of economies of the '*Hong Kong*' type has been rising fairly steadily since the nineties. The number of economies of the '*U.S.*' type has risen throughout the period, though less steadily.